

Advanced Algorithms (IV)

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Review

We learnt the Markov inequality

$$\Pr[X \geq a] \leq \frac{\mathbf{E}[X]}{a}$$

We can choose an increasing function f so that

$$\Pr[X \geq a] = \Pr[f(X) \geq f(a)] \leq \frac{\mathbf{E}[f(X)]}{f(a)}$$

$f(x) = x^2$ yields the Chebyshev's inequality

$$\Pr[|X - \mathbf{E}[X]| \geq a] \leq \frac{\mathbf{Var}[X]}{a^2} = \frac{\mathbf{E}[X^2] - \mathbf{E}[X]^2}{a^2}$$

What is a good choice of f ?

- f grows fast
- $\mathbf{E}[f(X)]$ is bounded and easy to calculate

Moment Generating Function

The function $f(x) = e^{tx}$ is a natural choice

The function $\mathbf{E}[f(X)] = \mathbf{E}[e^{tX}]$ is called the *moment generating function*

In some cases, $\mathbf{E}[e^{tX}]$ is easy to calculate...

Chernoff Bound

Assume $X = \sum_{i=1}^n X_i$, where each $X_i \sim \text{Ber}(p_i)$ is an independent Bernoulli variable with mean p_i

$$\begin{aligned}\mathbf{E}[e^{tX}] &= \mathbf{E}[e^{t \sum_{i=1}^n X_i}] \\ &= \prod_{i=1}^n \mathbf{E}[e^{X_i}] = \prod_{i=1}^n (p_i \cdot e^t + 1 - p_i) \\ &= \prod_{i=1}^n e^{p_i(e^t - 1)} = e^{\mathbf{E}[X](e^t - 1)}\end{aligned}$$

$$\text{Let } \mu = \mathbf{E}[X] = \sum_{i=1}^n p_i$$

For $t > 0$, we can deduce

$$\begin{aligned} \Pr[X > (1 + \delta)\mu] &= \Pr[e^{tX} \geq e^{t(1+\delta)\mu}] \\ &\leq \frac{\mathbf{E}[e^{tX}]}{e^{t(1+\delta)\mu}} = \frac{e^{(e^t-1)\mu} \star}{e^{t(1+\delta)\mu}} \end{aligned}$$

In order to obtain a tight bound, we optimize t to minimize \star

Since $\frac{e^{(e^t-1)\mu}}{e^{t(1+\delta)\mu}} = e^{\mu(e^t-1-t(1+\delta))}$, we can choose $t = \log(1 + \delta) > 0$.

$$\text{So } \Pr[X > (1 + \delta)\mu] \leq \left(\frac{e^\delta}{(1 + \delta)^{1+\delta}} \right)^\mu$$

We can similarly obtain (using $t < 0$)

$$\Pr[X < (1 - \delta)\mu] \leq \left(\frac{e^{-\delta}}{(1 - \delta)^{1-\delta}} \right)^\mu$$

To summarize, for $X = \sum_{i=1}^n X_i$, we have

- $\Pr[X \geq (1 + \delta)\mu] \leq \left(\frac{e^\delta}{(1 + \delta)^{1+\delta}} \right)^\mu$

- $\Pr[X \leq (1 - \delta)\mu] \leq \left(\frac{e^{-\delta}}{(1 - \delta)^{1-\delta}} \right)^\mu$

A more useful expression is that for $0 < \delta \leq 1$

- $\Pr[X \geq (1 + \delta)\mu] \leq e^{-\mu\delta^2/3}$

- $\Pr[X \leq (1 - \delta)\mu] \leq e^{-\mu\delta^2/2}$

Max Load

Recall in the max load problem, we throw n balls into n bins

The number of balls in i -th bin, $X_i \sim \text{Bin} \left(n, \frac{1}{n} \right)$

Note that $\mathbf{E}[X_i] = 1$, what is the probability that $X_i > \frac{c \log n}{\log \log n}$?

In this case, $1 + \delta = \frac{c \log n}{\log \log n}$.

Applying Chernoff bound, we obtain

$$\Pr[X_i \geq \frac{c \log n}{\log \log n}] \leq \frac{e^\delta}{(1 + \delta)^{1+\delta}} \leq n^{-c+o(1)},$$

which is tight in order comparing to our analytic result.

The Chernoff bound has a few drawbacks:

- each X_i needs to be *independent*.
- X_i is required to follow the $\text{Ber}(p_i)$

We will try to generalize the Chernoff bound to overcome these issues

Hoeffding Inequality

The Hoeffding Inequality generalizes to those X_i with $\mathbf{E}[X_i] = 0$ and $a_i \leq X_i \leq b_i$.

$$\Pr \left[\sum_{i=1}^n X_i \geq t \right] \leq \exp \left(- \frac{2t^2}{\sum_{i=1}^n (b_i - a_i)^2} \right)$$

The key property to establish Hoeffding inequality is an upper bound on the moment generating function

Lemma

Assume X satisfies $X \in [a, b]$ and $\mathbf{E}[X] = 0$, then

$$\mathbf{E}[e^{tX}] \leq \exp\left(\frac{t^2}{8}(b-a)^2\right)$$

You can find the proof of the lemma and Hoeffding inequality in the book *Probability and Computing*

Multi-Armed Bandit

In the problem of MAB, there are k bandits

- each bandit has a *unknown random* reward distribution f_i on $[0, 1]$ with $\mu_i = \mathbf{E}[f_i]$
- each round one can pull an arm i and obtain a reward $r \sim f_i$

The goal is to identify the best arm via trials

Regret of MAB

We assume $\mu_1 = \max_{1 \leq i \leq k} \mu_i$

If the game is played for T rounds, the best reward one can obtain is $T\mu_1$ *in expectation*

We are often not so lucky to achieve this, so the goal is to find a strategy to minimize

$$R(T) = T\mu_1 - \sum_{t=1}^T \mu_{a_t}$$

Regret Best Reward a_t - the arm actually pulled at round t

What is a good strategy?

We view $R(T)$ as a function of T and consider $T \rightarrow \infty$

If we eventually find the best arm, then $R(T) = o(T)$

If we fail to find the best arm, we will suffer a regret $\Omega(\Delta T)$, where Δ the gap between the optimal and suboptimal rewards

So we need the failure probability is $O(1/T)$

The Upper Confidence Bound Algorithm

We collect information up to round T

- $n_i(T)$ - number of times that i -th arm has been pulled
- $\hat{\mu}_i(T)$ - estimate of the mean μ_i , which is equal to
$$\frac{\sum_{t=1}^T \mathbf{1}[a_t = i] \cdot r(t)}{n_i(T)}$$
 if $n_i(T) \neq 0$ and $r(t)$ is the reward at t -th round

Choose the Best Arm So Far?

The most straightforward idea is to choose the arm with best $\hat{\mu}_i(T)$

The strategy might be inferior in case that we are unlucky so that the best arm performs bad at the first few trials.

So we have to add some offset term for those arms that are not “well-explored”

The UCB algorithm chooses the arm with largest

$$\hat{\mu}_i(T) + c_i(T)$$

$c_i(T)$ - the confidence
term of arm i at round T

Intuitively, $c_i(T)$ should be decreasing in n_i , so we give more chances to arms that have not been well-tested

Let's find out how to set $c_i(T)$

We need the following event to happen whp

$$\forall 2 \leq i \leq k, \quad \hat{\mu}_1(T) + c_1(T) > \hat{\mu}_i(T) + c_i(T)$$

A sufficient condition for this is

$$\hat{\mu}_1(T) + c_1(T) \boxed{>} \mu_1 \boxed{>} \mu_i + 2c_i(T) \boxed{>} \hat{\mu}_i(T) + c_i(T)$$

$$\boxed{>} + \boxed{>} : \forall j, \quad \hat{\mu}_j(T) - c_j(T) < \mu_j < \hat{\mu}_j(T) + c_j(T)$$

$$\boxed{>} : \forall i \geq 2, \quad c_i(T) < \frac{\mu_1 - \mu_i}{2}$$

Trade-off on $c_j(T)$

We apply Hoeffding inequality to bound the probability of

$$\forall j, \forall t \leq T, \quad \hat{\mu}_j(t) - c_j(t) < \mu_j < \hat{\mu}_j(t) + c_j(t)$$

$$\Pr[|\hat{\mu}_j(t) - \mu_j| > c_j(t)] \leq 2 \exp\left(-\frac{2c_j^2}{n_j(1/n_j)^2}\right) = 2 \exp(-2c_j^2 n_j)$$

So the Hoeffding inequality suggests us to choose

$$c_j(T) = \Omega\left(\sqrt{\frac{\log T}{n_j(T)}}\right)$$

For this choice of $c_i(T)$, the condition $c_i(T) < \frac{\mu_1 - \mu_i}{2}$ becomes to

$$\sqrt{\frac{c \log T}{n_i(T)}} < \frac{\mu_1 - \mu_i}{2}$$

This means $n_i(T) = \Omega(\log T)$, so we need to try each arm $\Omega(\log T)$ times for free!

Some tedious calculations are required to obtain the final regret bound, which is $\Theta(\log T)$