Algorithms for Big Data (V)

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Review of the Last Lecture

Last time, we learnt Misra-Gries and Count Sketch for Frequency Estimation.

The later has the advantage of being a linear sketch.

It also generalize to turnstile model.

Count Sketch

Algorithm Count Sketch

Init:

An array C[j] for $j \in [k]$ where $k = \frac{3}{\epsilon^2}$.

A random Hash function $h : [n] \to [k]$ from a 2-universal family.

A random Hash function $g:[n] \to \{-1,1\}$ from a 2-universal family.

On Input (y, Δ) :

 $C[h(y)] \leftarrow C[h(y)] + \Delta \cdot g(y)$

Output: On query a:

Output $\hat{f}_{\alpha} = g(\alpha) \cdot C[h(\alpha)]$.

The Performance

We can apply the median trick to obtain:

- ▶ $\Pr\left[\left|\widehat{f}_{\alpha} f_{\alpha}\right| \geqslant \varepsilon \|\mathbf{f}\|_{2}\right] \leqslant \delta;$
- ▶ it costs $O\left(\frac{1}{\epsilon^2}\log\frac{1}{\delta}\left(\log m + \log n\right)\right)$ bits of memeory.

Today we will see another simple sketch algorithm.

Count-Min

We assume that for each entry (y, Δ) , it holds that $\Delta \geqslant 0$.

Algorithm Count-Min

Init:

An array C[i][j] for $i \in [t]$ and $j \in [k]$ where $t = log(1/\delta)$ and $k = 2/\epsilon$.

Choose t independent random Hash function $h_1, \ldots, h_t : [n] \to [k]$ from a 2-universal family.

On Input (y, Δ) :

For each $i \in [t]$, $C[i][h_i(y)] \leftarrow C[i][h_i(y)] + \Delta$.

Output: On query a:

Output $\hat{f}_{\alpha} = \min_{1 \leq i \leq t} C[i][h(\alpha)].$

Analysis

Obviously we have $f_{\alpha} \leqslant \widehat{f}_{\alpha}$.

Our algorithm overestimates only if for some $b \neq a$, $h_i(b) = h_i(a)$. Let $Y_{i,b}$ be the indicator of this event.

Let X_i be $C[i][h_i(a)]$. Then

$$\mathbf{E}\left[\widehat{X}_i\right] = \sum_{b \in [n]} f_b \mathbf{E}\left[Y_{i,b}\right] = f_\alpha + \sum_{b \in [n]: b \neq \alpha} f_b \mathbf{E}\left[Y_{i,b}\right] = f_\alpha + \frac{\sum_{b \in [n]: b \neq \alpha} f_b}{k} \leqslant f_\alpha + \frac{\|\mathbf{f}\|_1}{k}.$$

Thus,

$$\Pr[|X_{\mathfrak{i}} - f_{\mathfrak{a}}| \geqslant \varepsilon ||\mathbf{f}||_{1}] \leqslant \frac{||\mathbf{f}||_{1}}{||\mathbf{f}||_{1}} = \frac{1}{2}.$$

Since our output is the minimum out of t independent X_i 's,

$$\begin{split} \mathbf{Pr}\left[\widehat{\mathbf{f}}_{\alpha} - \mathbf{f}_{\alpha} \geqslant \epsilon \|\mathbf{f}\|_{1}\right] &= \mathbf{Pr}\left[|\min\left\{X_{1}, \ldots, X_{t}\right\} - \mathbf{f}_{\alpha}| \geqslant \|\mathbf{f}\|_{1}\right] \\ &= \mathbf{Pr}\left[\bigwedge_{i=1}^{t}\left(|X_{i} - \mathbf{f}_{\alpha}| \geqslant \epsilon \|\mathbf{f}\|_{1}\right)\right] \\ &= \prod_{i=1}^{t} \mathbf{Pr}\left[|X_{i} - \mathbf{f}_{\alpha}| \geqslant \epsilon \|\mathbf{f}\|_{1}\right] \leqslant 2^{-t} = \delta. \end{split}$$

The algorithm computes a linear sketch using

$$O\left(\frac{1}{\varepsilon}\log\frac{1}{\delta}\cdot(\log m + \log n)\right)$$

bits of memory.

It can be generalized to turnstile model (Exercise).

Frequency Moments

The k-th frequency moment of a stream is

$$F_k \triangleq \sum_{\mathbf{j} \in [n]} f_{\mathbf{j}}^k = \|\mathbf{f}\|_k^k.$$

For example, F_2 is the size of self-join of a relation r.

Many problems we met before can be viewed as estimating $F_{\mathbf{k}}$ for some special \mathbf{k} .

AMS Estimator for F_{k}

Given $(\alpha_1, \ldots, \alpha_m)$, then algorithm first sample a uniform index $I \in [m]$.

It then count the number of entries a_i with $a_i = a_I$ and $j \ge J$.

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Algorithm AMS Estimator for F<sub>k</sub>
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Init: (m, r, a) \leftarrow (0, 0, 0).
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On Input (y, Δ) :

$$m \leftarrow m + 1, \ \beta \sim Ber(\frac{1}{m});$$

if
$$\beta = 1$$
 then

$$a \leftarrow y, r \leftarrow 0;$$

end if

if
$$y = a$$
 then $r \leftarrow r + 1$ end if

Output:

$$m\left(r^k-(r-1)^k\right).$$

Analysis

We first compute the expectation of the output X.

Assuming a = i at the end of algorithm, then

$$\mathbf{E}\left[X\mid\alpha=j\right] = \mathbf{E}\left[m(r^k - (r-1)^k)\mid\alpha=j\right] = \sum_{i=1}^{f_j} \frac{1}{f_i} \cdot m\left(i^k - (i-1)^k\right) = \frac{m}{f_i} \cdot f_j^k.$$

Therefore,

$$\mathbf{E}[X] = \sum_{i=1}^{n} \mathbf{Pr}[\alpha = j] \cdot \mathbf{E}[X \mid \alpha = j] = \sum_{i=1}^{n} \frac{f_j}{m} \cdot \frac{m}{f_j} \cdot f_j^k = F_k.$$

The Variance

$$\begin{aligned} & \text{Var}\left[X\right] \leqslant \mathbf{E}\left[X^{2}\right] = \sum_{j=1}^{n} \frac{f_{j}}{m} \sum_{i=1}^{f_{j}} \frac{1}{f_{j}} \cdot m^{2} \left(i^{k} - (i-1)^{k}\right)^{2} \\ & \leqslant m \sum_{j=1}^{n} \sum_{i=1}^{f_{j}} k i^{k-1} \left(i^{k} - (i-1)^{k}\right) \\ & \leqslant m \sum_{j=1}^{n} k f_{j}^{k-1} \sum_{i=1}^{f_{j}} \left(i^{k} - (i-1)^{k}\right) \\ & = m \sum_{j=1}^{n} k f_{j}^{k-1} \cdot f_{j}^{k} = k \left(\sum_{j=1}^{n} f_{j}\right) \left(\sum_{j=1}^{n} f_{j}^{2k-1}\right). \end{aligned}$$

Assume $k \ge 1$ and let $f_* \triangleq \max_{i \in [n]} f_i$.

$$\begin{aligned} \mathbf{Var}\left[X\right] &\leqslant k \sum_{j=1}^{n} f_{j} \cdot \left(f_{*}^{k-1} \sum_{j=1}^{n} f_{j}^{k}\right) \\ &\leqslant k \sum_{j=1}^{n} f_{j} \cdot \left(\left(f_{*}^{k}\right)^{\frac{k-1}{k}} \sum_{j=1}^{n} f_{j}^{k}\right) \\ &\leqslant k \sum_{j=1}^{n} f_{j} \cdot \left(\sum_{j=1}^{n} f_{j}^{k}\right)^{\frac{k-1}{k}} \sum_{j=1}^{n} f_{j}^{k} \end{aligned}$$

Applying Jensen's inequality on
$$g(z) = z^{1/k}$$
, we can bound also

$$\leqslant k\sum_{j=1}f_j\cdot\left(\sum_{j=1}f_j^k\right)\sum_{j=1}f_j^k$$
 Applying Jensen's inequality on $g(z)=z^{1/k}$, we can bound above by

 $k \sum_{i=1}^{n} (f_{j}^{k})^{\frac{1}{k}} \left(\sum_{i=1}^{n} f_{j}^{k} \right)^{\frac{k-1}{k}} \sum_{i=1}^{n} f_{j}^{k} \leqslant kn^{1-1/k} \left(\sum_{i=1}^{n} f_{j}^{k} \right)^{\frac{1}{k}} \left(\sum_{i=1}^{n} f_{j}^{k} \right)^{\frac{k-1}{k}} \sum_{i=1}^{n} f_{j}^{k} = kn^{1-1/k} F_{k}^{2}.$

Therefore,

$$\Pr\left[|X - F_k| \geqslant \varepsilon F_k\right] \leqslant \frac{kn^{1-1/k}}{\varepsilon^2}.$$

Now we can apply the standard averaging trick and median trick.

To kill the $n^{1-1/k}$ factor in the variance, we need to average $\Omega\left(n^{1-1/k}\right)$ estimates.

An (ε, δ) estimator requires

$$O\left(\frac{1}{\epsilon^2}\log\frac{1}{\delta}kn^{1-1/k}\left(\log m + \log n\right)\right)$$

bits of memory.

The Tug-of-War Sketch

The following simple algorithm for F_2 outperforms AMS by using only $O(\log n + \log m)$ bits.

Algorithm Tug-of-War Sketch

Init:

A random Hash function $h:[n] \to \{-1,1\}$ from a 4-universal family.

 $\boldsymbol{x} \leftarrow 0.$

On Input (y, Δ) :

 $x \leftarrow x + \Delta \cdot h(y)$

Output:

Output x^2 .

Analysis

Let X be the value of x at the end of our algorithm.

$$\mathbf{E}\left[X^{2}\right] = \mathbf{E}\left[\left(\sum_{\mathbf{j}\in[\mathbf{n}]}f_{\mathbf{j}}h(\mathbf{j})\right)^{2}\right] = \mathbf{E}\left[\sum_{\mathbf{j}\in[\mathbf{n}]}f_{\mathbf{j}}^{2}h(\mathbf{j})^{2} + \sum_{\mathbf{i},\mathbf{j}\in[\mathbf{n}]:\mathbf{i}\neq\mathbf{j}}f_{\mathbf{i}}f_{\mathbf{j}}h(\mathbf{i})h(\mathbf{j})\right] = F_{2}.$$

Using the property of 4-universal Hash family, we have

$$\begin{split} \mathbf{E}\left[X^{4}\right] &= \sum_{i,j,k,\ell \in [n]} f_{i}f_{j}f_{k}f_{\ell}\mathbf{E}\left[h(i)h(j)h(k)h(\ell)\right] \\ &= \sum_{j \in [n]} f_{j}^{4}\mathbf{E}\left[h(j)^{4}\right] + 6\sum_{i,j \in [n]: j > i} f_{i}^{2}f_{j}^{2}\mathbf{E}\left[h(i)^{2}h(j)^{2}\right] = F_{4} + 6\sum_{i,j \in [n]: j > i} f_{i}^{2}f_{j}^{2}. \end{split}$$

Therefore

$$\mathbf{Var} [X^{2}] = \mathbf{E} [X^{4}] - (\mathbf{E} [X^{2}])^{2}$$

$$= F_{4} - F_{2}^{2} + 6 \sum_{i,j \in [n]: j > i} f_{i}^{2} f_{j}^{2}$$

$$= F_{4} - F_{2}^{2} + 3(F_{2}^{2} - F_{4}) \leq 2F_{2}^{2}.$$

Finally, we apply the median trick and it costs

$$O\left(\frac{1}{\varepsilon^2}\log\frac{1}{\delta}\left(\log n + \log m\right)\right)$$

bits of memory.